Estimating new structural equation models with the Bayesian methods (future ideas for Mplus and time series modeling)

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- Overview of time series models and how to run some time series models in Mplus
- Overview of time series models and SEM, dynamic factor analysis (DFA) models and Kalman filter
- How to incorporate DFA and Kalman filter in Mplus with Bayes estimation

Overview of time series models

- Time series models: analysis of data measured at successive time instants
- Wide and solid applications: econometrics, signal processing, and mathematical finance
- Intensive longitudinal data (ILD) in the social sciences: more longitudinal data are collected that makes very frequent observations using new tools for data collection such as palm pilots, smart phones etc. Walls & Schafer (2006)
- Jahng S., Wood, P. K., & Trull, T. J., (2008) Analysis of Affective Instability in Ecological Momentary Assessment: Indices Using Successive Difference and Group Comparison via Multilevel Modeling. Psychological Methods, 13, 354-375
- Measurement instrument for a "mood" factor collected several times a day for several months
- Ecological Momentary Assessment (EMA): involves repeated sampling of subjects current behaviors and experiences in real time, in subjects natural environments

- New book: Bolger & Laurenceau (2012), Intensive Longitudinal Methods: An Introduction to Diary and Experience Sampling Research New York: Guilford Press
- Uses Mplus prominently and includes examples with Mplus inputs

- Time series models are models for the disturbances/residuals in a model as a function of time
- It is easy to combine with SEM for any longitudinal data. Two separate models: a **Structural model** and **Disturbances model** = **Time series model**
- Any residual variable can be modeled further (beyond the SEM part) with a time series model
- Modeling frameworks that combine SEM and Time-series: Dynamic Factor Analysis and Kalman Filter model

Table : Google search results in millions

Term	Pages
Time series	22.8
Factor analysis	4.3
Principal component analysis	4.3
Kalman filter	3.7
Mplus	1.7
Multiple imputation	0.25
Dynamic factor analysis	0.09

- We generate data on 100 individuals with 200 time points
- Using a Kalman filter model with 5 indicators 1 factor
- Two plots of within level factor, using AR(1)
- Using autoregressive coefficient of 0.95 and 0
- Both plots represent zero mean within factor value with the same variance (disturbance process) (the between factor is not included at time *t* or any growth/trend values)
- Four colors represent 4 individuals/clusters

Residual plot with autocorrelation of 0



Residual plot with autocorrelation of 0.95



- Without time series modeling we are assuming picture 1 where the reality maybe picture 2
- Without time series modeling we fail to understand the process behind the data
- Ignoring the correlations between the residuals will underestimate the SE
- The predictive power of a model without the time series model will be worse

• AR(1) - autoregressive model with lag 1, with autocorrelation ρ

$$\varepsilon_t = \rho \varepsilon_{t-1} + \xi_t$$

• For
$$t > 1$$
, $Var(\xi_t) = \sigma^2$

- For the first term which is not regressed on anything, $Var(\xi_1) = \sigma^2/(1-\rho^2)$
- Stationary process, i.e., the variance of the disturbance is constant across time $Var(\varepsilon_t) = \sigma^2/(1-\rho^2)$

• implies a correlation matrix for ε_t , t = 1, ..., T

$$\begin{pmatrix} 1 & \rho & \rho^2 & \rho^3 & \dots & \rho^T \\ \rho & 1 & \rho & \rho^2 & \dots & \rho^{T-1} \\ \dots & \dots & \dots & \dots & \dots & \dots \\ \rho^T & \rho^{T-1} & \rho^{T-2} & \rho^{T-3} & \dots & 1 \end{pmatrix}$$

• $Corr(\varepsilon_t, \varepsilon_s) = \rho^{t-s}$

Disturbance models - AR(1). How to run in Mplus - Method 1 as in UG example 6.17

Disturbance models - AR(1). How to run in Mplus - Method 2

MODEL: y2-y4 pon y1-y3 (corr); y2-y4 (v); y1 (v1); MODEL CONSTRAINT: v1=v/(1-corr*corr);

Disturbance models - AR(1). Comparison: Method 1 v.s. Method 2

Loglikelihood				
HO Va	alue		-6682.537	
MODEL RESULTS				
	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
New/Additional	l Parameters			
CORR	0.718	0.010	<mark>68.81</mark> 6	0.000
Loglikelihood				
HO Va	alue		-6682.537	
MODEL RESULTS				
	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
¥2 ON	CONTRACTOR NO.	10122353	The second second	Market South
¥1	0.718	0.010	68.816	0.000
ты	omin Aspanouhov, Mplus	Malue 1	5/ 28	

```
MONTECARLO:
     NAMES = y1-y100;
     NOBS = 100;
     NREPS = 100;
MODEL POPULATION:
y2-y100 pon y1-y99*0.7 (corr);
y2-y100*1 (v); y1*1.96 (v1);
[v1-v100*0] (m);
MODEL:
y2-y100 pon y1-y99*0.7 (corr);
y2-y100*1 (v); y1*1.96 (v1);
[y1-y100*0] (m);
MODEL CONSTRAINT:
v1=v/(1-corr*corr);
```

Disturbance models - AR(1). Simulation study results, T=100

MODEL RESULTS

			ESTIMATES		S.E.	M. S. E.	95% % Sig
		Population	Average	Std. Dev.	Average		Cover Coeff
¥2	ON						
¥1		0.700	0.7000	0.0072	0.0071	0.0001	0.940 1.000
Means							
Y1		0.000	0.0012	0.0098	0.0100	0.0001	0.970 0.030
Varianc	es						
¥1		1.960	1.9632	0.0441	0.0470	0.0019	0.970 1.000
¥2		1.000	1.0009	0.0131	0.0142	0.0002	0.980 1.000
Y3		1.000	1.0009	0.0131	0.0142	0.0002	0.980 1.000

Disturbance models - AR(1). Limitations of multivariate setup

- It works with T=200, with T=300 or more not computable due to memory problems.
- If the model is multivariate then maximum possible T will have to be divided by P (the number of variables)
- If T is large (T=10000 in the next example) the model can be run as a univariate model where the data is organized in a long format with the lag 1 variables set as covariate *Y*1. Ignore the initial equation.
- For large T ignoring the initial equation has negligible effect on the estimation.
- It is important to set *Y*1 as covariate to preserve log-likelihood value

Disturbance models - AR(1). Univariate setup Method 3: duplication

Data setup

0.20510475	-0.378137
-0.590985813	0.20510475
-1.027777453	-0.590985813
-2.863705363	-1.027777453
1.674698659	-2.863705363
0.788097665	1.674698659
0.766737416	0.788097665
2.031961354	0.766737416
2.394893339	2.031961354
0.050893335	2.394893339
-0.095748666	0.050893335
1.094087833	-0.095748666

Disturbance models - AR(1). Univariate setup Method 3: duplication

variable: names = y y1; data: file = dbl.dat; model: y on y1*0.25;

Disturbance models - AR(1). Univariate setup Method 3: results

MODEL	RESULTS				
					Two-Tailed
		Estimate	S.E.	Est./S.E.	P-Value
Y	ON				
Y	1	0.269	0.010	27.873	0.000
Inte	rcepts				
Y		-0.002	0.010	-0.189	0.850
Resi	dual Varian	ces			
Y		1.026	0.015	70.700	0.000

• The AR(1) is a model for the disturbances. How do you setup the model in the presence of Y predictors such as latent variables from a growth model or observed covariates

Disturbance models - AR(1) for growth model Method 1

i s | y100 y201 y302 y403; MODEL: y1-y4 (resvar); y1-y3 PWITH y2-y4 (p1); y1-y2 PWITH y3-y4 (p2); y1 WITH y4 (p3); MODEL CONSTRAINT: NEW (corr); p1 = resvar*corr; p2 = resvar*corr**2; p3 = resvar*corr**3;

Disturbance models - AR(1) for growth model Method 2

```
MODEL:
           i s | y100 y201 y302 y403;
     v1-v4@0;
     f1 by y101;
     f2 by y201;
     f3 by y301;
     f4 by y4@1;
     f2-f4 pon f1-f3 (corr);
     f2-f4 (v); f1 (v1);
MODEL CONSTRAINT:
v1=v/(1-corr*corr);
```

Disturbance models - more models

• MA(1) - moving average model with lag 1

$$\varepsilon_t = \xi_t + \theta \xi_{t-1}$$

 $Cov(\varepsilon_t, \varepsilon_s) = 0$ if t - s > 1

• ARMA(1,1) - autoregressive model with lag 1 and moving average model with lag 1

$$\varepsilon_t = \rho \varepsilon_{t-1} + \xi_t + \theta \xi_{t-1}$$

• ARMA(p,q) - autoregressive model with lag 1 and moving average model with lag 1

$$\varepsilon_t = \rho_1 \varepsilon_{t-1} + \ldots + \rho_p \varepsilon_{t-p} + \xi_t + \theta_1 \xi_{t-1} + \ldots + \theta_q \xi_{t-q}$$

• ARMA(p,q) models have been very successful disturbance models in practice

```
MODEL:
y2-y10 pon y1-y9*0.7 (corr);
[y1-y10*0] (m);
y2-y10@0;
y1*1.04 (v1);
f0-f10 with f0-f10@0;
f0-f10*1(v);
f0 by y1*0.3 (theta);
f1 by y1@1 y2*0.3 (theta);
f2 by y2@1 y3*0.3 (theta);
. . . . . .
f9 by y9@1 y10*0.3 (theta);
f10 by v10@1;
```

```
MODEL CONSTRAINT:
v1=corr*corr*v*(1+theta*theta)/(1-corr*corr);
```

DFA and Kalman filter

- Kalman filter≈ DFA ≈ ARMA disturbance model for the factor over time.
- Molenaar (1985). A dynamic factor model for the analysis of multivariate time series. Psychometrika
- Zhang, Hamaker and Nesselroade (2008) Comparisons of Four Methods for Estimating a Dynamic Factor Model. Structural Equation Modeling
- Zhang and Nesselroade (2007) Bayesian Estimation of Categorical Dynamic Factor Models
- Justiniano (2004). Estimation and model selection in dynamic factor analysis. PhD dissertation

• Direct autoregressive factor score (DAFS)

 $Y_t = \Lambda f_t + \varepsilon_t$ $f_t = \rho f_{t-1} + \xi_t$

• White noise factor score (WNFS)

$$Y_t = \Lambda f_t + \Lambda_1 f_{t-1} + \varepsilon_t$$

DAFS model estimation in Mplus following Molenaar (1985): duplication

- Suppose that we have 3 indicator 1 factor DAFS model
- Organize the data as follows: duplication Y1 Y2 Y3 YLag11 YLag12 YLag13 1 2 3 4 5 6 7 8 9 1 2 3 10 12 12 7 8 9

```
. . . .
```

DAFS model estimation in Mplus following Molenaar (1985): duplication

• write the model as

f by Y1-Y3 (l1-l3); f@1; flag1 by YLag11-YLag13 (l1-l3); flag1@1; [Y1-Y3](m1-m3); [YLag11-YLag13](m1-m3); Y1-Y3(v1-v3); YLag11-YLag13(v1-v3); f on flag1 (rho);

• This is not the original suggestion by Molenaar but the above setup yields the same likelihood

DAFS model estimation in Mplus following Molenaar (1985): duplication

- Problem 1: observations are not independent
- Problem 2: tests of fit are meaningless
- Problem 3: in AR(1) regression equation only observed data from lag 1 have effect
- Problem 4: Standard error would have to be adjusted
- Problem 5: Not true ML
- Problem 6: simulations show that it works well in some cases but no guarantee that will work well in all cases

• Consider the model

$$Y_{ti} = Y_{B,i} + \Lambda f_{ti} + \varepsilon_{ti}$$
$$f_{ti} = f_{B,i} + f_{W,ti}$$
$$f_{W,ti} = \rho f_{W,t-1,i} + \xi_{ti}$$

for i=1,...,T and i=1,...,N

- This model has a two-level factor structure with individually specific factor mean and AR(1) structure for the disturbances
- This model is more general than DAFS and WNFS because it adds two-level structure
- One way to think about this model is that the two-level part explains the mean of f_{ti} within a cluster while the AR(1) explains the covariance within a cluster and abandons the assumption that observations are conditionally independent given the between level random effects

If we ignore the AR(1) part of the model, Mplus MCMC will take the following steps to generate parameters and latent variable

• Step 1.

$$[f_{B,i}, Y_{B,i} | Y_{ti}, parameters]$$

multivariate normal

• New Step 1

 $[f_{B,i}, Y_{B,i}|f_{W,ti}, Y_{ti}, parameters]$

conditioning on $f_{W,ti}$ makes this a feasible step otherwise the Y_{ti} between level random effects are not independent, loss of efficiency

- Step 2. $[f_{w,it}|f_{B,i}, Y_{ti}, parameters]$ simply subtract the between parts and conduct independent factor analysis update for each $f_{w,it}$
- New Step 2. [*f_{w,it}*|*f_{B,i}*, *Y_{ti}*, *parameters*] has to be broken into *T* steps

 $[f_{w,i1}|f_{w,i,-1}f_{B,i}, Y_{ti}, parameters] \\ [f_{w,i2}|f_{w,i,-2}f_{B,i}, Y_{ti}, parameters]$

 $[f_{w,iT}|f_{w,i,-T}f_{B,i}, Y_{ti}, parameters]$

where $f_{w,i,-t}$ means all within level factors but $f_{w,i,t}$, i.e., $f_{w,i,1}$,... $f_{w,i,t-1}f_{w,i,t+1}$,... $f_{w,i,T}$

• New Step 2.

$$[f_{w,it}|f_{w,i,-t}f_{B,i},Y_{ti}, parameters]$$

is based on the following equations

 $Y_{W,ti} = \Lambda f_{W,ti} + \varepsilon_{ti}$ $f_{W,t+1,i} = \rho f_{W,ti} + \xi_{t+1,i}$ $f_{W,ti} = \rho f_{W,t-1,i} + \xi_{ti}$

- The factor that we need to update has one more indicator $f_{W,t+1,i}$ and a predictor $f_{W,t-1,i}$
- For t=1 and t=T the model looses one of the two new equations. It is important to properly account for the ending equations probably for T < 200.
- For t=1, $f_{W,ti} = \sigma^2/(1-\rho^2)$. It is important to generate $f_{W,1i}$ with this bigger variance otherwise the Baysian generation of $f_{W,ti}$ will not look stationary and would have an increasing variance.

The estimation of slopes, intercepts, and loadings, given the latent variables is identical, including *ρ*, however the estimation of *ρ* is based on one fewer equations than all other parameters:
 T – 1 equations, which actually causes change in the use of sufficient statistics

$$f_{W,ti} = \rho f_{W,t-1,i} + \xi_{ti}$$

- The estimation of the variance parameters is the same with one exception: $Var(\xi_{ti})$. For t > 1, $Var(\xi_{ti}) = \sigma^2$ but for t = 1 $Var(\xi_{ti}) = \sigma^2/(1 - \rho^2)$. To resolve this problem in updating σ^2 we use $\sqrt{(1 - \rho^2)}\xi_{1i}$ instead of ξ_{1i} .
- The convergence will be fast as long as the factor is measured well by the indicators, that would make all variables nearly observed.

%within% f by y1-y5; f on f%1; ! AR(1) disturbance model %between% fb by y1-y5;

- Developing disturbance models in Mplus will combine SEM framework with the time series framework and deliver many new modeling possibilities
- Consider time intense data Y_{it} analyzed as cross-classified, crossed by individual and time. The variance decomposition is

$$Var(Y_{it}) = S_i + S_t + S_{it}$$

and now we would be able to add autoregressive structure on the within level as well.